

# Module 8: Interval Estimation

- $\theta$  is fixed while  $\theta_{\text{hat}_n}$  is a random variable which provides the single best value to estimate  $\theta$
- $\theta_{\text{hat}}$  is unbiased when  $\text{bias} = E(\theta_{\text{hat}_n}) - \theta = 0$
- $\theta_{\text{hat}}$  is consistent when  $\theta_{\text{hat}_n} \rightarrow \theta$
- Mean squared error  $\text{MSE} = E(\theta_{\text{hat}_n} - \theta)^2 = \text{bias}(\theta_{\text{hat}_n}) + V(\theta_{\text{hat}_n})$
- If bias  $\rightarrow 0$  and se  $\rightarrow 0$  as  $n \rightarrow \infty$  then  $\theta_{\text{hat}_n}$  is consistent
- Probability is stronger than samples, probability standard error eventually converges to 0 as  $n$  approaches infinity but samples converge to a normal distribution which is not necessarily the same as the population distribution.
- **We use the estimator variability (se) to provide an interval of parameter values that are "supported" by the sample.**

A  $1 - \alpha$  confidence interval for a parameter  $\theta$  is an interval  $C_n = (a; b)$  where  $a = a(X_1, \dots, X_n)$  and  $b = b(X_1, \dots, X_n)$  are functions of the data such that:  $P(\theta \in C_n) \geq 1 - \alpha$ ; Where  $\theta$  is the actual population mean.  $C_n$  is random and  $\theta$  is fixed.

The confidence interval  $(a; b)$  capture the true mean with confidence  $1 - \alpha$ . We commonly use 95% confidence intervals which corresponds to  $\alpha = .05$ . This does **NOT** mean there is  $1 - \alpha$  chance/probability the parameter falls in the interval. The correct interpretation: If we repeatedly take samples of size  $n$  from a fixed and stable population and build a 95% confidence intervals, 95% of these intervals would contain the true unknown parameter.

## CI For Mean of a Normal Distribution

If  $\sigma^2$  is known:  $X_{\text{bar}} \pm Z_{\alpha/2} * \sigma_x$

If  $\sigma^2$  is unknown:  $X_{\text{bar}} \pm t_{(\alpha/2, n-1)} * S / \sqrt{n}$ ; Where  $S^2 = 1/(n-1) * \sum(x_i - x_{\text{bar}})^2$

Using  $S$  in place of SD causes more uncertainty, thus increasing the size of the CI.

We can similarly find the confidence interval of a proportion in a similar manner:

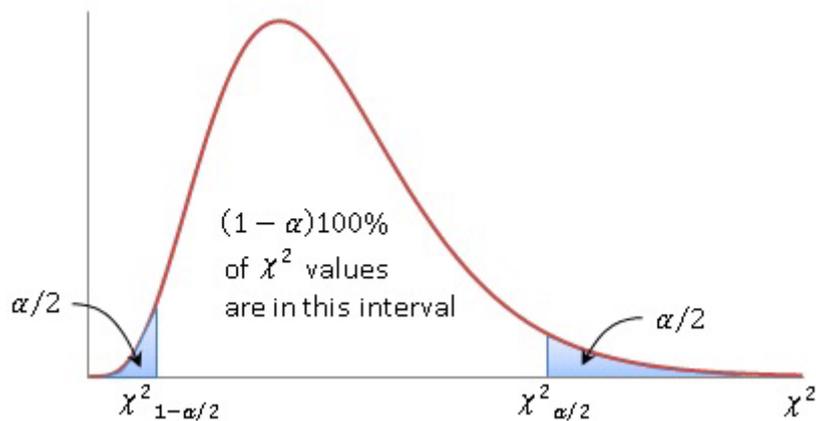
$$C_n = \left( \hat{p}_n - z_{\alpha/2} \sqrt{\frac{\hat{p}_n(1 - \hat{p}_n)}{n}}, \hat{p}_n + z_{\alpha/2} \sqrt{\frac{\hat{p}_n(1 - \hat{p}_n)}{n}} \right)$$

## Chi-Square Distribution

$$Q = \sum_{i=1}^n X_i^2$$

The above represents the chi-squared distribution with  $n$  degrees of freedom.  $E(Q) = n$  and  $V(Q) = 2n$ .

The distribution of  $\chi^2_{n-1}$  is not symmetrical, so instead of centering our CI (a,b) on the mean, we look for symmetry so that the bounds  $P(\theta < a) = \alpha/2$  and  $P(\theta > b) = \alpha/2$ .



We derive the variance of a distribution through Fisher's theorem (not shown). The CI comes out to:

$$\mathbb{P}(\sigma^2 < a) = \alpha/2 \Rightarrow \mathbb{P}\left(\frac{(n-1)S^2}{\sigma^2} > \frac{(n-1)S^2}{a}\right) = \alpha/2 \quad (1)$$

$$\mathbb{P}\left(\frac{(n-1)S^2}{\sigma^2} > \chi^2_{(\alpha/2, n-1)}\right) = \alpha/2 \quad (2)$$

• Thus  $a = \frac{(n-1)S^2}{\chi^2_{(\alpha/2, n-1)}} \leq \sigma^2 \leq b = \frac{(n-1)S^2}{\chi^2_{(1-\alpha/2, n-1)}}$

Pearson's product Moment Correlation Coefficient is between -1 and 1 and represents the correlation between 2 variables.

$$r = \frac{\sum(x - \bar{x})(y - \bar{y})}{\sqrt{\sum(x - \bar{x})^2 \sum(y - \bar{y})^2}}$$

Although rarely used, you could find a confidence interval for this value.

$$Z = \frac{1}{2} \log\left(\frac{1+r}{1-r}\right)$$

$$se(Z) = \frac{1}{\sqrt{n-3}}$$

$$\frac{\exp(2F) - 1}{\exp(2F) + 1} \text{ to } \frac{\exp(2G) - 1}{\exp(2G) + 1}$$

$$F = Z - \frac{Z_{\alpha/2}}{\sqrt{n-3}} \text{ and } G = Z + \frac{Z_{\alpha/2}}{\sqrt{n-3}}$$

---

Revision #7

Created 24 August 2022 14:06:18 by Elkip

Updated 24 August 2022 15:51:56 by Elkip